

Leonard G. Johnson, EDITOR

Vol. 1

May, 1971

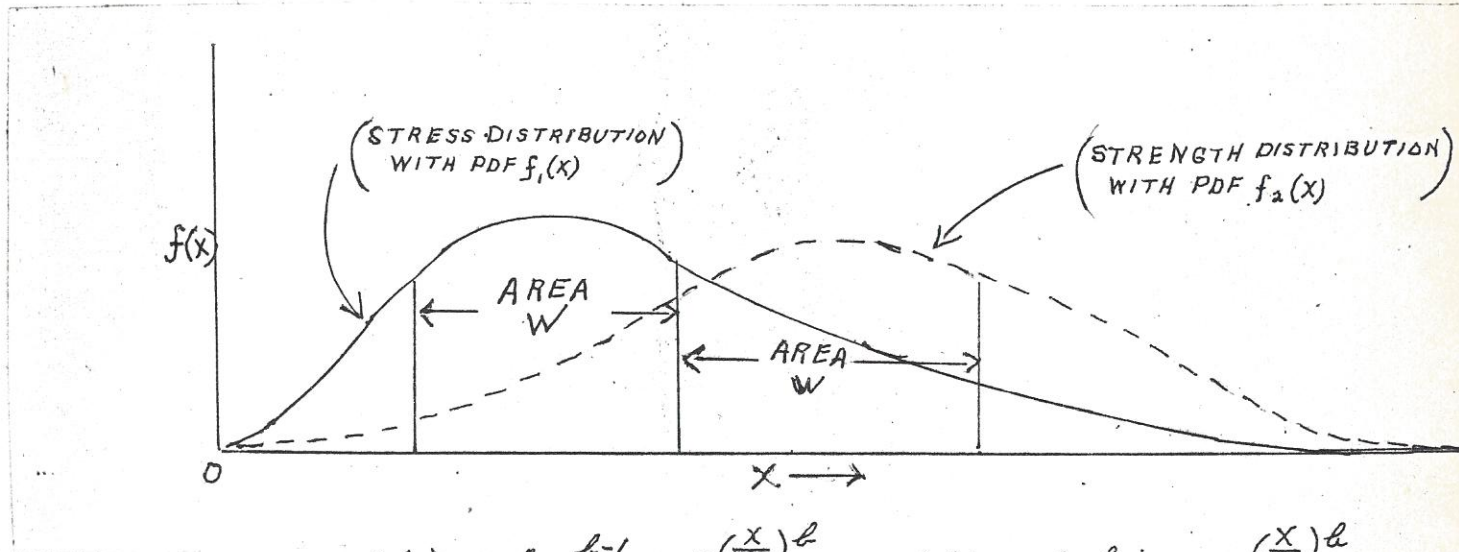
Bulletin 2

Page 1

As a start for this second bulletin we shall derive the last formula given in bulletin No. 1, i.e.

$$P = \text{Probability that strength} > \text{stress} = \frac{\text{LOG}\left(\frac{1-W}{2}\right)}{\text{LOG}\left(\frac{1-W^2}{4}\right)}$$

The formula is obtained as follows:



We assume $f_1(x) = \frac{bx^{b-1}}{\theta_1^b} e^{-\left(\frac{x}{\theta_1}\right)^b}$ AND $f_2(x) = \frac{bx^{b-1}}{\theta_2^b} e^{-\left(\frac{x}{\theta_2}\right)^b}$,

Which amounts to assuming stress and strength have 2-parameter Weibull distributions of the same slope b , but different characteristic values.

Now, prob. ($x \leq \text{stress} \leq x+dx$) = $f_1(x) dx$

prob. [(strength > stress) if ($x \leq \text{stress} \leq x+dx$)] = $1-F_2(x)$

where $F_2(x) = 1 - e^{-\left(\frac{x}{\theta_2}\right)^b}$

Prob. [(strength > stress) and ($x \leq \text{stress} \leq x+dx$)]

$$= [1-F_2(x)] f_1(x) dx$$

$$= e^{-\left(\frac{x}{\theta_2}\right)^b} \cdot \frac{bx^{b-1}}{\theta_1^b} e^{-\left(\frac{x}{\theta_1}\right)^b} dx$$

$$= \frac{bx^{b-1}}{\theta_1^b} e^{-x^b \left(\frac{1}{\theta_1^b} + \frac{1}{\theta_2^b}\right)} dx$$

$$\therefore \text{Prob. (strength > stress)} = \int_0^{\infty} \frac{bx^{b-1}}{\theta_1^b} e^{-x^b \left(\frac{1}{\theta_1^b} + \frac{1}{\theta_2^b}\right)} dx$$

$$= \frac{\frac{1}{\theta_1^b}}{\left(\frac{1}{\theta_1^b} + \frac{1}{\theta_2^b}\right)} \int_0^{\infty} bx^{b-1} \left(\frac{1}{\theta_1^b} + \frac{1}{\theta_2^b}\right) e^{-x^b \left(\frac{1}{\theta_1^b} + \frac{1}{\theta_2^b}\right)} dx$$

$$= \frac{1}{1 + \left(\frac{\theta_1}{\theta_2}\right)^b} \left[-e^{-x^b \left(\frac{1}{\theta_1^b} + \frac{1}{\theta_2^b}\right)} \right]_0^{\infty} = \frac{1}{1 + \left(\frac{\theta_1}{\theta_2}\right)^b}$$

Since central bands of width W just touch, it follows that

$$B \left(\frac{1+W}{2} \right) \text{ Level of Stress} = B \left(\frac{1-W}{2} \right) \text{ Level of Strength}$$

$$\text{or } \theta_1 \left(\ln \frac{1}{1 - \frac{1+W}{2}} \right)^{\frac{1}{b}} = \theta_2 \left(\ln \frac{1}{1 - \frac{1-W}{2}} \right)^{\frac{1}{b}}$$

$$\text{or } \theta_1^b \ln \left(\frac{2}{1-W} \right) = \theta_2^b \ln \left(\frac{2}{1+W} \right)$$

$$\text{or } \left(\frac{\theta_1}{\theta_2} \right)^b = \frac{\ln \left(\frac{2}{1+W} \right)}{\ln \left(\frac{2}{1-W} \right)} = \frac{\ln \left(\frac{1+W}{2} \right)}{\ln \left(\frac{1-W}{2} \right)}$$

$$\begin{aligned} \text{PROB. (strength} > \text{stress)} &= \frac{1}{1 + \left(\frac{\theta_1}{\theta_2} \right)^b} = \frac{1}{1 + \frac{\ln \left(\frac{1+W}{2} \right)}{\ln \left(\frac{1-W}{2} \right)}} \\ &= \frac{\ln \left(\frac{1-W}{2} \right)}{\ln \left(\frac{1-W}{2} \right) + \ln \left(\frac{1+W}{2} \right)} \\ &= \frac{\ln \left(\frac{1-W}{2} \right)}{\ln \left(\frac{1-W^2}{4} \right)} \end{aligned}$$

Q. E. D.

THE USE OF THE "ENTROPY" CONCEPT IN RELIABILITY

Let $R(x)$ = the reliability of an item to target X .

Then we define

$$\text{"Entropy" at } X = G(x) = \ln \frac{1}{R(x)} \quad (1)$$

We call this an "entropy" because it involves the logarithm of a probability, namely, the negative logarithm of the survival probability $R(x)$. Another name for this quantity $G(x)$ would be the "cumulative hazard". Thus, our definition of "entropy at X " is the same as the "cumulative hazard at X ".

$$\text{From (1): } R(x) = e^{-G(x)} \quad (2)$$

or, in words,

$$\text{Reliability to target } x = e^{-\text{(Entropy at } x\text{)}}$$

Now, let $F(x)$ = Probability of failure of the item before target x
(i.e. $F(x)$ = CDF of X)

$$\text{Then } F(x) = 1 - R(x) = 1 - e^{-G(x)} \quad (3)$$

From (3) it can be seen that a 2-parameter Weibull distribution $F(x) = 1 - e^{-\left(\frac{x}{\theta}\right)^{\beta}}$ is a special case in which "entropy at x " is given by $G(x) = \left(\frac{x}{\theta}\right)^{\beta}$

PROPERTIES OF "ENTROPY"

PROPERTY 1: Since life has a CDF $F(x) = 1 - e^{-\epsilon(x)}$ it follows that "entropy" has a CDF

$$F(\epsilon) = 1 - e^{-\epsilon}$$

Thus, "entropy" is always exponentially distributed with a Mean Value of Unity. This property of entropy can be used to estimate the additional life of an item which has survived to some time X_0 .

Let X_0 = time which an item has survived thus far.

Let $\epsilon(X_0)$ = entropy at X_0

Since the average additional entropy to failure is 1, we add 1 more unit of entropy to the entropy at X_0 , and solve for the X which has a total entropy of $\epsilon(X_0) + 1$.

$$\text{Thus, } \epsilon(X) = \epsilon(X_0) + 1$$

For example, suppose an item has survived 100 hours and has a Weibull CDF with $b = 2$ and $\theta = 150$ hours. The entropy thus far is $(\frac{100}{150})^2 =$

$$\frac{4}{9} = \epsilon(X_0); (X_0 = 100) \text{ add 1 more unit of entropy: } \frac{4}{9} + 1 = \frac{13}{9}$$

Now find a larger X such that $(\frac{x}{150})^2 = \frac{13}{9}$

$$\therefore \frac{x}{150} = \sqrt{\frac{13}{9}} = \frac{1}{3} \sqrt{13} = \frac{3.60553}{3} = 1.20185$$

$$x = 150 (1.20185) = 180.3 \text{ hrs.}$$

Thus, an additional 80.3 hrs. will be expired before the average additional entropy of 1 will be reached.

PROPERTY II

If an item has survived to time X_0 , at which time the entropy is $G(X_0)$, we can predict with confidence C that the item will survive at least to a total time X such that

$$G(x) = G(x_0) + \ln \frac{1}{C}$$

For example, in the previous problem where

$$\left. \begin{array}{l} X_0 = 100 \text{ hrs.} \\ b = 2 \\ \theta = 150 \text{ hrs.} \end{array} \right\}$$

We can say with confidence 50% that the total life will at least be X where

$$\left(\frac{x}{150} \right)^2 = \left(\frac{100}{150} \right)^2 + \ln \frac{1}{.5} = \frac{4}{9} + .69315 = 1.03759$$

$$\therefore \frac{x}{150} = \sqrt{1.03759} = 1.0186$$

$$x = 150 (1.0186) = 152.8 \text{ hrs.}$$

with 50% confidence, the item will run at least another 52.8 hours

PARAMETER ESTIMATION USING ENTROPY

Problem: Given n failures at X_1, X_2, \dots, X_n and k suspended items at U_1, U_2, \dots, U_k from a 2-parameter Weibull Distribution of slope b . Estimate the value of θ using the fact that the average entropy to failure is 1.

$$\text{Entropy at } X_1 = G(x_1) = \left(\frac{x_1}{\theta} \right)^b$$

$$\text{Entropy at } X_2 = G(x_2) = \left(\frac{x_2}{\theta} \right)^b$$

$$\vdots \quad \quad \quad \vdots \quad \quad \quad \vdots$$

$$\text{Entropy at } X_N = e^{-\left(\frac{x_N}{\theta}\right)^b}$$

$$\text{Entropy at } U_1 = e^{-\left(\frac{u_1}{\theta}\right)^b}$$

$$\text{Entropy at } U_2 = e^{-\left(\frac{u_2}{\theta}\right)^b}$$

$$\text{Entropy at } U_k = e^{-\left(\frac{u_k}{\theta}\right)^b}$$

$$\text{Sum of entropies} = \left(\frac{x_1}{\theta}\right)^b + \left(\frac{x_2}{\theta}\right)^b + \dots + \left(\frac{x_N}{\theta}\right)^b + \left(\frac{u_1}{\theta}\right)^b + \left(\frac{u_2}{\theta}\right)^b + \dots + \left(\frac{u_k}{\theta}\right)^b$$

Since entropies to failure are exponentially distributed with an average of unity, it follows that

$$\text{Ave. entropy to failure} = \frac{\text{sum of entropies thus far}}{\text{total failures thus far}} = 1$$

$$\text{Thus, } \frac{\left(\frac{x_1}{\theta}\right)^b + \left(\frac{x_2}{\theta}\right)^b + \dots + \left(\frac{x_N}{\theta}\right)^b + \left(\frac{u_1}{\theta}\right)^b + \left(\frac{u_2}{\theta}\right)^b + \dots + \left(\frac{u_k}{\theta}\right)^b}{n} = 1$$

$$\text{or } \theta^b = \frac{x_1^b + x_2^b + \dots + x_N^b + u_1^b + u_2^b + \dots + u_k^b}{n}$$

$$\text{Thus, } \theta = \left(\frac{x_1^b + x_2^b + \dots + x_N^b + u_1^b + u_2^b + \dots + u_k^b}{n} \right)^{\frac{1}{b}}$$

This estimate of the characteristic life θ is the same as the so-called maximum likelihood estimate of θ .